

Econ 501 Final — December 12 2022

You are allowed one piece of letter size paper with writing on one side only. Show your work throughout.

1. Suppose that $\{\mathbf{x}_i\}$ are i.i.d. draws from a distribution with density function

$$f(x) = 2 \max(x, 0) \exp(-x^2/\theta_0)/\theta_0,$$

for some $\theta_0 > 0$.

- (a) Derive the asymptotic distribution of the maximum likelihood estimator of θ_0 . Express the asymptotic variance in terms of θ_0 only.

Answer: $N(0, \theta_0^2)$

- (b) Propose a test for the hypothesis $H_0 : \theta_0 = 1$ versus $H_1 : \theta_0 \neq 1$ at a 5% significance level. Please be complete in your answer, i.e. include a decision rule, a test statistic, and a critical value.

Answer: several possibilities, but an ordinary t test with $\mathbf{t} = n^{-1} \sum_{i=1}^n (\mathbf{x}_i^2 - 1)$ does fine. The critical value is then 1.96.

- (c) Show that your test is consistent.

Answer: like in class

- (d) Determine the local power of your test if the local alternative is $\theta_0 = 1 + \lambda/\sqrt{n}$.

Answer: like in class

2. Let

$$\begin{bmatrix} \mathbf{x} \\ \mathbf{y} \end{bmatrix} \sim N\left(0, \begin{bmatrix} 1 & \rho^\top \\ \rho & I \end{bmatrix}\right),$$

for some vector ρ whose elements are all between -1 and 1. Express $\mathbb{V}(\mathbf{y} | \mathbf{x} \leq x)$ in terms of ρ , x and the inverse Mills ratio $m(x) = \phi(x)/\Phi(x)$ where ϕ, Φ are the standard normal density function and standard normal distribution function, respectively. [2]

Answer: Note that $\mathbb{E}(\mathbf{x} | \mathbf{x} \leq x) = -m(x)$, $\mathbb{E}(\mathbf{x}^2 | \mathbf{x} \leq x) = 1 - xm(x)$, $\mathbb{E}(\mathbf{y} | \mathbf{x} \leq x) = -xm(x)\rho$, $\mathbb{E}(\mathbf{y}\mathbf{y}^\top | \mathbf{x} \leq x) = I - xm(x)\rho\rho^\top$. Thus, the answer is $I - m(x)\{x + m(x)\}\rho\rho^\top$

3. Suppose that $2n(\mathbf{x}_n - 1) \xrightarrow{d} N(0, 1)$ and $\mathbf{y}_n \xrightarrow{p} 2$. Derive the limit distribution of $\{n(\mathbf{x}_n^2 - 1)\}^{\mathbf{y}_n}$.
4. Recall that the Hoeffding inequality says that if $\{\mathbf{x}_i\}$ is independent (but not necessarily identically distributed) and for some numbers $\{a_i\}$, $\{b_i\}$, $\mathbb{P}(a_i \leq \mathbf{x}_i \leq b_i) = 1$ then

$$\forall z > 0 : \mathbb{P}\left(\left|\sum_{i=1}^n (\mathbf{x}_i - \mathbb{E}\mathbf{x}_i)\right| > z\right) \leq 2 \exp\left(\frac{-2z^2}{\sum_{i=1}^n (b_i - a_i)^2}\right).$$

Suppose that the support of the i.i.d. $\{\mathbf{x}_i\}$ is $[-1/\sqrt{2}, 1/\sqrt{2}]$. Construct a test for the hypothesis $H_0 : \mu_0 = 0$ against the alternative hypothesis $H_1 : \mu_0 \neq 0$ where $\mu_0 = \mathbb{E}\mathbf{x}_1$ that

has rejection probability no greater than α under H_0 . Use a decision rule of the form: 'reject H_0 if $|\bar{\mathbf{x}}| > C_\alpha$ '

Answer: Reject H_0 if $|\bar{\mathbf{x}}| > \sqrt{n^{-1} \log(2/\alpha)}$.

5. True or false? Please provide a justification if true and a counter example if false.

(a) $\forall x, y : F_{\mathbf{x}|\mathbf{y}}(x|y)F_{\mathbf{y}}(y) = F_{\mathbf{xy}}(x, y)$, where the $F_{\mathbf{x}|\mathbf{y}}$ denotes a conditional distribution function, $F_{\mathbf{xy}}$ a joint distribution function and $F_{\mathbf{y}}$ a marginal distribution function.

Answer: False generally: only true if there is independence. If $f(x, y) = 2\mathbb{1}(0 \leq x \leq y \leq 1)$ then $F(x|y) = \mathbb{1}(x \leq y) \min(x, y)/y$, $F(y) = y^2$, and $F(x, y) = (2xy - x^2)\mathbb{1}(x \leq y)$ for $0 \leq x, y \leq 1$.

(b)

$$\forall u, v : \frac{\mathbb{E} \exp\{i(u\mathbf{x} + v\mathbf{y})\}}{\mathbb{E} \exp(iu\mathbf{x})\mathbb{E} \exp(iv\mathbf{y})} = 1 \Rightarrow \mathbf{x}, \mathbf{y} \text{ are independent.}$$

Answer: true: if the joint characteristic function is the product of the marginal characteristic functions then there is independence

(c) $\mathbf{x}$ is conditionally mean independent of $\mathbf{y} \Rightarrow \mathbf{y}$ is conditionally mean independent of $\mathbf{x}$.

Answer: False. Take $\mathbf{y} = \mathbf{x}^2$ where $\mathbf{x}$ has an even distribution. Then $\mathbb{E}(\mathbf{x} | \mathbf{y} = y) = 0$ for all y , but $\mathbb{E}(\mathbf{y} | \mathbf{x} = x) = x^2$.

6. Suppose that $\theta_0 = \arg \min_{\theta \in \Theta} \Omega(\theta)$ (so θ_0 is unique) where Ω is continuous and Θ compact. Suppose further that $\sup_{\theta \in \Theta} |\hat{\Omega}(\theta) - \Omega(\theta)| = o_p(1)$. Prove that $\hat{\theta} = \arg \min_{\theta \in \Theta} \hat{\Omega}(\theta) \xrightarrow{p} \theta_0$.

Answer: notes